

Biostat/Stat 576
Statistical Methods for Survival Analysis

Lecture 2
April 2, 2009

Chapter 1. Introduction

1. A brief history of time-to-event analysis
2. Time-to-event and censoring
3. Life tables
4. Counting processes and martingales

4. Counting Processes and Martingales

- T : time-to-event, failure time, survival time, nonnegative random variable
- Characterization
 1. $f(t)$: density function
 2. $F(t) = \int_0^t f(u)du$: (cumulative) distribution function, (cumulative) risk
 3. $S(t) = \bar{F}(t) = 1 - F(t)$: survival function
 4. $\Lambda(t) = -\log S(t)$: (cumulative) hazard function, (cumulative) risk
 5. $\lambda(t) = \Lambda'(t)$: hazard function, (instantaneous) risk
 6. $m(t) = E(T - t \mid T \geq t)$: mean residual life function

- Center piece: hazard function

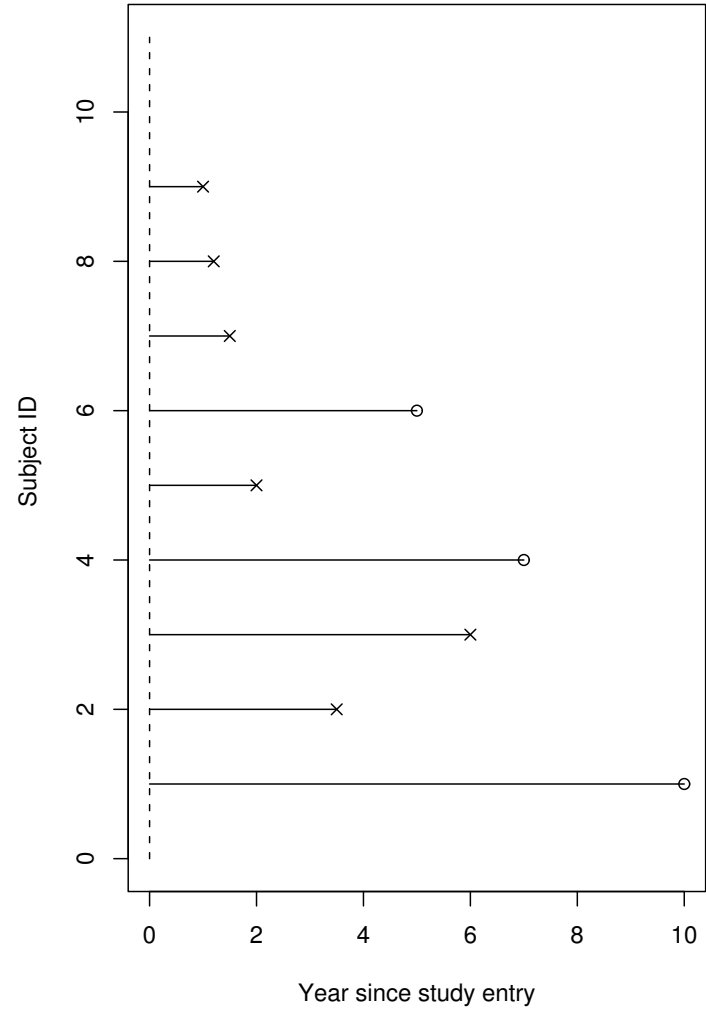
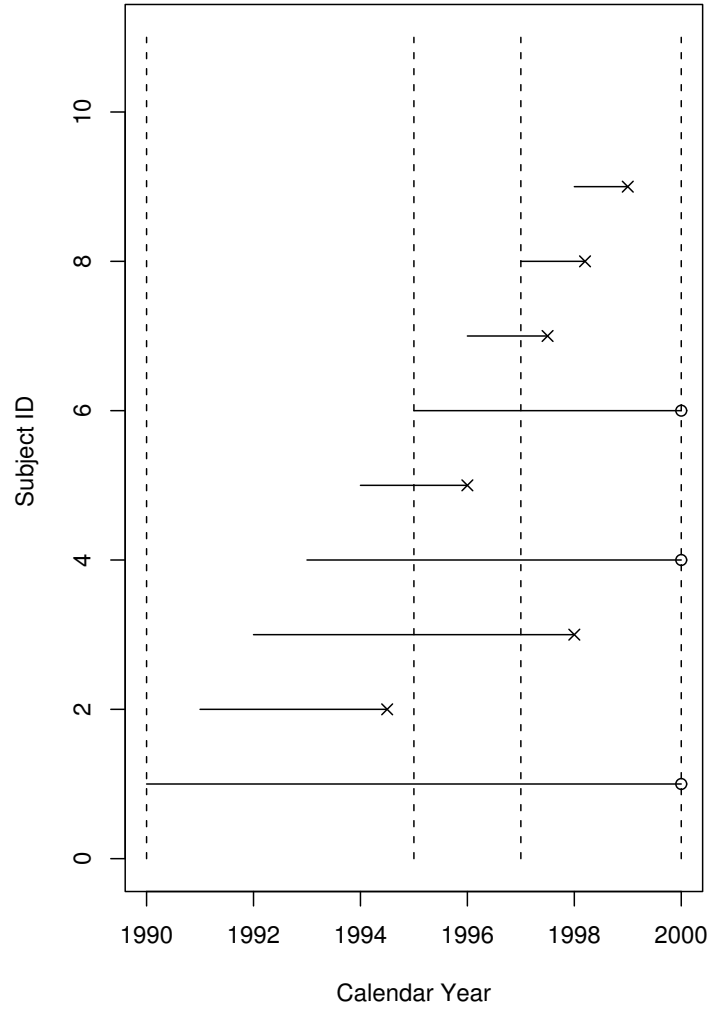
1. $\lambda(t) = \lim_{\Delta t \rightarrow 0} \Pr\{t \leq T \leq t + \Delta t \mid T \geq t\} / \Delta t$

2. $\lambda(t)\Delta t \approx \Pr\{t \leq T \leq t + \Delta t \mid T \geq t\}$

3. $\lambda(t) = f(t)/S(t) = -S'(t)/S(t) = [-\log S(t)]'_t$

4. $\lambda(t) = [1 + m'(t)]/m(t)$

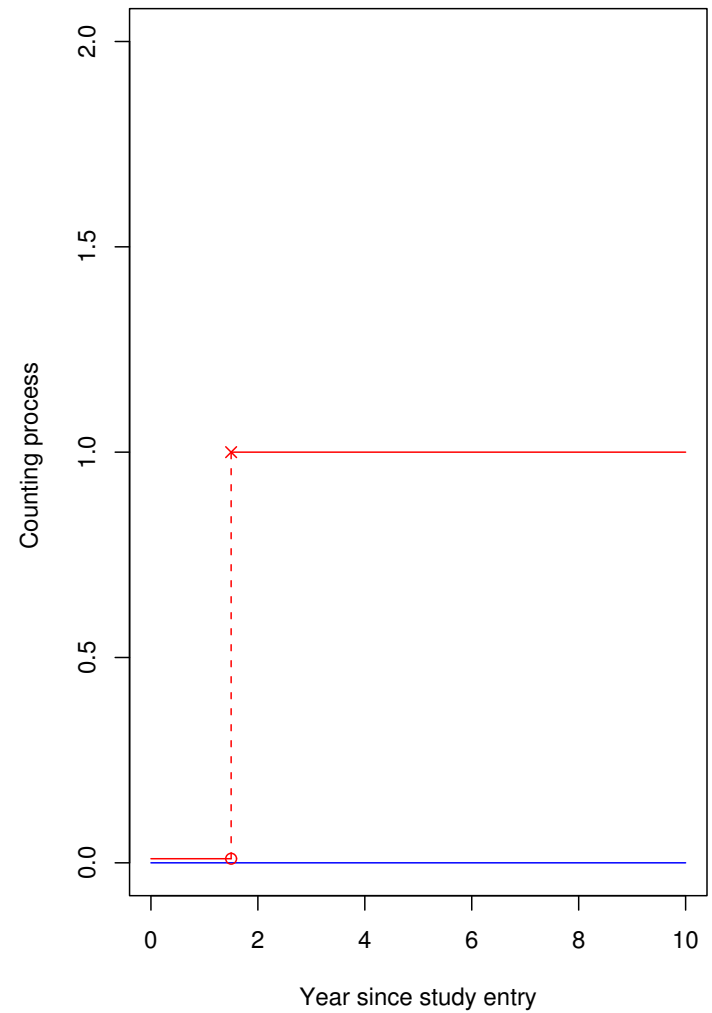
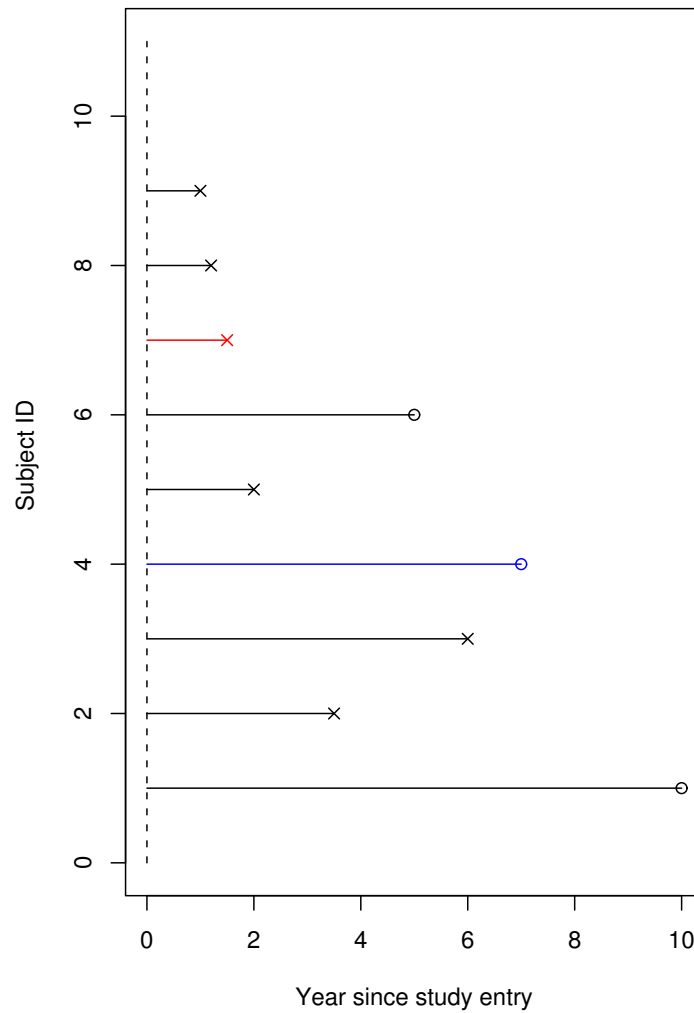
- Typical observed right-censored data



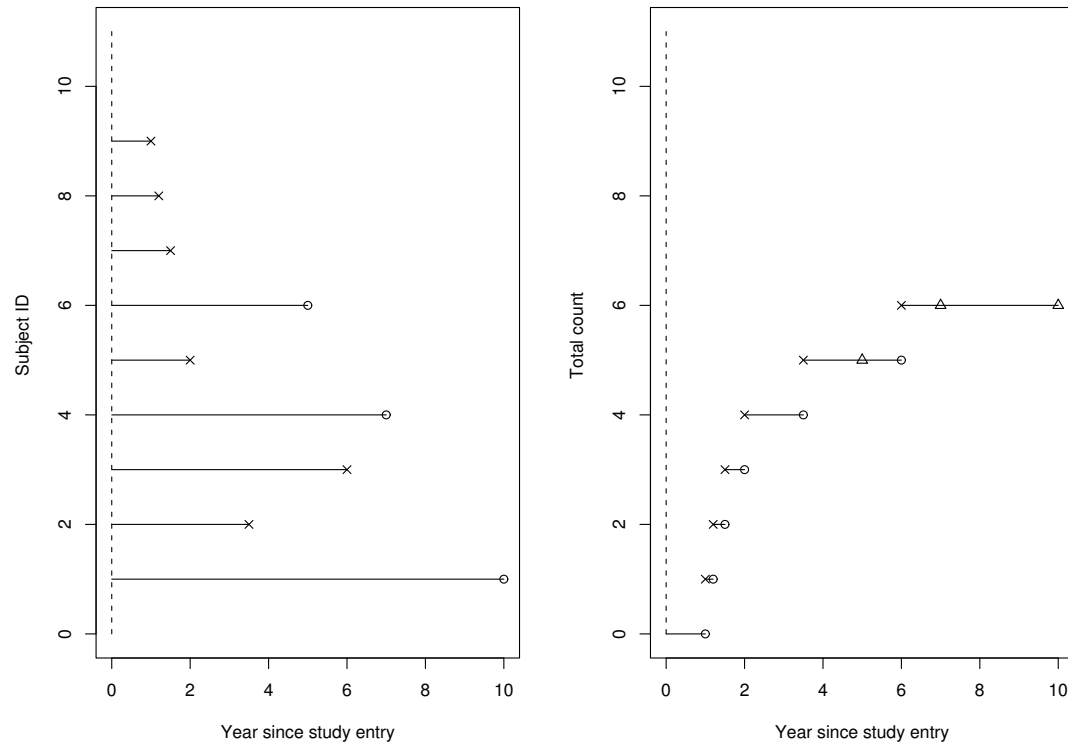
- Data formulation
 1. T : time-to-event
 2. C : censoring time, usually right-censored assumed
 3. $X = \min(T, C)$: (censored) survival time
 4. $\Delta = I(T \leq C)$: event/censoring indicator
 5. Z or $Z(t)$: (time-varying) covariates
 6. $i = 1, 2, \dots, n$: subjects
 7. $\{(X_i, \Delta_i, Z_i), i = 1, 2, \dots, n\}$: collected data
- Major goals of survival analysis
 1. one-sample: identify T 's characterization
 2. two-sample: compare T 's characterizations
 3. regression: identify association between Z and T 's characterization

- Counting processes

1. $N_i(t) = I(X_i \leq t, \Delta_i = 1)$: count the event prior to time t

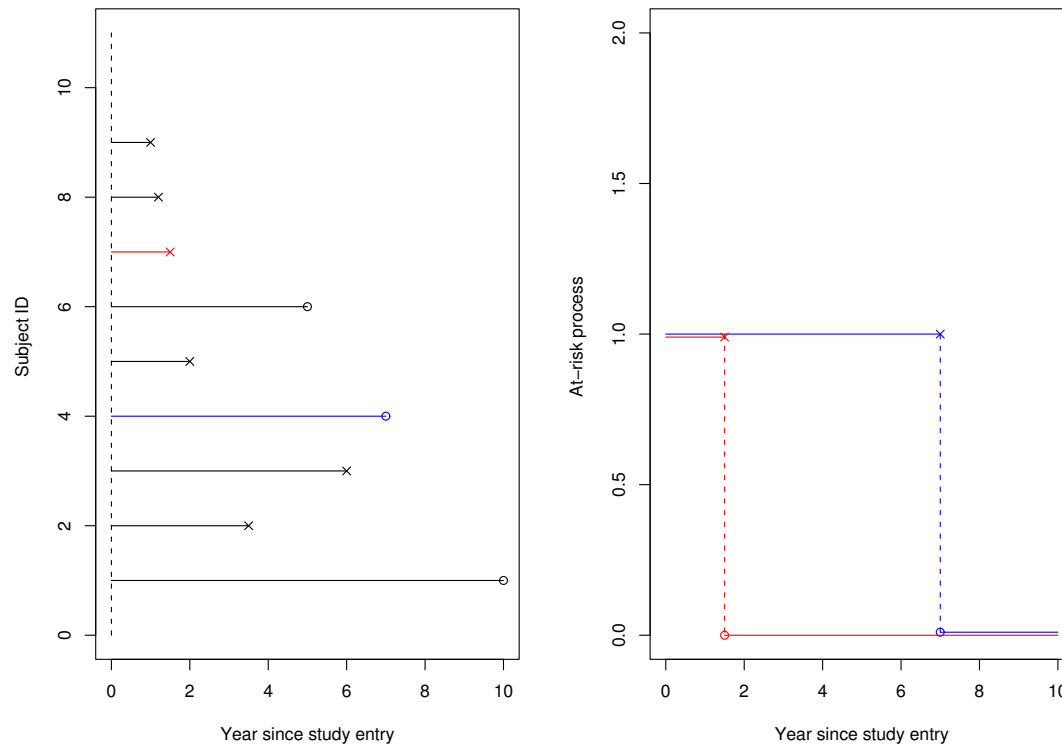


2. $N(t) = \sum_i N_i(t)$: total number of events prior to time t



- (a) stochastic processes
- (b) non-decreasing sample paths
- (c) step function taking jumps of size one
- (d) cadlag: right-continuous with left-hand limits (*continué à droit, limité à gauche*)

- At-risk process: $Y_i(t) = I(X_i \geq t)$



1. $Y(t) = \sum_i Y_i(t)$: total number of subjects at risk
2. right-continuous with right-hand limit
3. “predictable”

- Steiltjes integral: $\int_0^t f(u)dG(u)$
 1. $G(\cdot): [0, \infty) \mapsto \mathcal{R}$
 2. right-continuous, increasing and differentiable with derivative $g(\cdot)$ except at a countable set of points $\{0 < t_0 < t_1 < \dots\}$
 3. $\int_0^t f(u)dG(u) = \sum_{t_i \leq t} \int_{t_{i-1}}^{t_i} f(u)g(u)du + \sum_{t_i \leq t} f(t_i)[G(t_i) - G(t_i-)]$

- Stochastic integrals: Steiltjes integral of one stochastic process with respect to another

- Example

1. Data: $3 = (3, 1), 4^+ = (4, 0), 8 = (8, 1), 9^+ = (9, 0), 11 = (11, 1)$

2. Nelson-Aalen estimator

$$\hat{\Lambda}(t) = \int_0^t \frac{I\{Y(u) > 0\}}{Y(u)} dN(u)$$

t	$Y(t)$	$dN(t)$	$\hat{\Lambda}(t)$
$[0, 3)$	5	0	0
$[3]$	5	1	1/5
$(3, 4]$	4	0	1/5+0/4
$(4, 8)$	3	0	1/5+0/4+0/3
$[8]$	3	1	1/5+1/3
$(8, 9]$	2	0	1/5+1/3+0/2
$(9, 11)$	1	0	1/5+1/3+0/2+0/1
$[11]$	1	1	1/5+1/3+1/1
$(11, \infty)$	0	0	1/5+1/3+1/1

- What's so SPECIAL about time?
 1. progressive: data accumulation
 2. irreversible: uni-direction
 3. when things have happened, there is nothing we can do about them but learn from them!

- Filtration: \mathcal{F}_t
 1. entire historical information up to t
 2. data collected: $\{(N_i(t), Y_i(t), Z_i); i = 1, 2, \dots, n\}$
 3. $\mathcal{F}_t = \sigma\{(N_i(u), Y_i(u), Z_i); u \leq t, i = 1, 2, \dots, n\}$
 4. \mathcal{F}_{t-} : information accumulated right before t

- Martingale
 1. $\{M(t); t \geq 0\}$ is a stochastic process
 2. for any $s < t$, $E[M(t) | \mathcal{F}_s] = M(s)$
 3. $E[M(t) | \mathcal{F}_{t-}] = M(t-) \iff E[dM(t) | \mathcal{F}_{t-}] = 0$

- Example: calculate $E[dN(t) | \mathcal{F}_{t-}]$

1. $dN(t) = 0$ or 1

2. $E[dN(t) | \mathcal{F}_{t-}] = \Pr\{dN(t) = 1 | \mathcal{F}_{t-}\}$

3. if $Y(t) = 0 \implies Y(t-) = 0$, then $\Pr\{dN(t) = 1 | \mathcal{F}_{t-}\} = 0$

4. if $Y(t) = 1 \implies Y(t-) = 1$, then

$$\Pr\{dN(t) = 1 | \mathcal{F}_{t-}\} = \Pr\{t \leq X \leq t + dt, \Delta = 1 | X \geq t\}$$

$X = \min(T, C) \geq t \Leftrightarrow T \geq t, C \geq t$, if T and C are independent, then

$$\Pr\{dN(t) = 1 | \mathcal{F}_{t-}\} = \Pr\{t \leq T \leq t + dt | T \geq t\} = \lambda(t)dt$$

5. $E[dN(t) | \mathcal{F}_{t-}] = Y(t)\lambda(t)dt$

6. $M(t) = N(t) - \int_0^t Y(u)\lambda(u)du$ is a martingale

Chapter 2. Parametric methods

1. Parametric distributions
2. Likelihood functions and MLE

1. Parametric distributions

- Density function: $f(t) = \lim_{\Delta t \rightarrow 0} \Pr\{t \leq T \leq T + \Delta t\}$
 1. $\int_0^{\infty} f(u) du = 1$
 2. $F(t) = \int_0^t f(u) du, S(t) = \int_t^{\infty} f(u) du$
 3. $\mu(t) = \int_0^t S(u) du, \mu = ET = \int_0^{\infty} S(u) du$
 4. Denominator: all the subjects
- Hazard function: $\lambda(t) = \lim_{\Delta t \rightarrow 0} \Pr\{t \leq T \leq T + \Delta t \mid T \geq t\} / \Delta t$
 1. $\lambda(t) = f(t) / S(t)$
 2. Denominator: subjects survival up to t
- Mean residual life function: $m(t) = E(T - t \mid T \geq t)$
 1. Residual time: $R_T(t) = (T - t)I(T \geq t) = (T - t)^+ = (T - t) \vee 0$
 2. $m(t) = \int_t^{\infty} S(u) du / S(t)$

- Example: exponential distribution
 1. Density function: $f(t) = \lambda e^{-\lambda t}$, $S(t) = e^{-\lambda t}$
 2. Constant hazard function: $\lambda(t) = \lambda$, $\Lambda(t) = \lambda t$
 3. Constant mean residual life: $m(t) = \mu = 1/\lambda$
 4. Lack of memory: $f_{R_T(t)}(u) = f_T(u)$
 5. Coefficient of variation: s.d./mean=1, for judging relative dispersion
 6. For any T , $\Lambda(T)$ is exponential

- Other parametric distributions: $\{f(t; \theta); \theta \in \Theta\}$
 1. Gamma
 2. Weibull
 3. Log normal
 4. log logistic
 5. Ref: Kalbfleisch & Prentice, §§2.2.1-2.2.8

2. Likelihood functions and MLE

- Statistcal inferences
 1. dimension reduction: summarizing data by models with parameters
 2. parameter estimation: identifying parameters and providing error bound
 3. estimation efficiency: utilizing maximum information
 4. model checking: assessing model adequacy
 5. model-based prediction

- Likelihood inferences

1. choose a parametric family of distributions: $f(t | Z_i; \theta)$
2. if no censoring, $\{(T_i, Z_i); i = 1, 2, \dots, n\}$
 - likelihood function based on observed $T_i = t_i$ is

$$\mathcal{L}(\theta) = \prod_{i=1}^n f(t_i | Z_i; \theta)$$

3. if censored, survival data become $\{(X_i, \Delta_i, Z_i); i = 1, 2, \dots, n\}$
- likelihood contribution is $\Pr\{X_i = x_i, \Delta_i = \delta_i, Z_i = z_i\}$
 - $\Delta_i = 1 \implies X_i = t_i$, likelihood contribution is $f_T(t_i | Z_i; \theta)S_C(t_i | Z_i)$
 - $\Delta_i = 0 \implies X_i = c_i$, likelihood contribution is $f_C(c_i | Z_i)S_T(c_i | Z_i; \theta)$
 - likelihood function based on observed data is then

$$\mathcal{L}(\theta) = \prod_{i=1}^n f(x_i | Z_i; \theta)^{\delta_i} S(x_i | Z_i; \theta)^{1-\delta_i}$$

4. What are the key assumptions?
- Z_i do not have information about θ
 - conditional independence: given Z_i , T_i and C_i are independent
 - examples
 - * clinical trials
 - * Z_i : treatment assignment
 - * C_i : censoring due to side-effects

5. Parameter estimation: maximum likelihood estimation (MLE)

– likelihood function:

$$\begin{aligned}\mathcal{L}(\theta) &= \prod_{i=1}^n f(x_i | Z_i; \theta)^{\delta_i} S(x_i | Z_i; \theta)^{1-\delta_i} \\ &= \prod_{i=1}^n \lambda(x_i | Z_i; \theta)^{\delta_i} S(x_i | Z_i; \theta)\end{aligned}$$

– parameter estimation: $\hat{\theta} = \arg \max_{\theta} \mathcal{L}(\theta) = \arg \max_{\theta} l(\theta)$

$$l(\theta) = \log \mathcal{L}(\theta)$$

$$= \sum_{i=1}^n [\delta_i \log \lambda(x_i | Z_i; \theta) + \log S(x_i | Z_i; \theta)]$$

$$= \sum_{i=1}^n [\delta_i \log \lambda(x_i | Z_i; \theta) - \Lambda(x_i | Z_i; \theta)]$$

$$= \sum_{i=1}^n \left[\delta_i \log \lambda(x_i | Z_i; \theta) - \int_0^{x_i} \lambda(u | Z_i; \theta) du \right]$$

– score function

$$\begin{aligned}l'(\theta) &= \sum_{i=1}^n \left[\delta_i \frac{\lambda'_\theta(x_i | Z_i; \theta)}{\lambda(x_i | Z_i; \theta)} - \int_0^\infty I(x_i \geq u) \lambda'_\theta(u | Z_i; \theta) du \right] \\ &= \sum_{i=1}^n \left[\int_0^\infty \frac{\lambda'_\theta(u | Z_i; \theta)}{\lambda(u | Z_i; \theta)} dN_i(u) - \int_0^\infty Y_i(u) \lambda'_\theta(u | Z_i; \theta) du \right] \\ &= \sum_{i=1}^n \int_0^\infty \frac{\lambda'_\theta(u | Z_i; \theta)}{\lambda(u | Z_i; \theta)} [dN_i(u) - Y_i(u) \lambda(u | Z_i; \theta) du]\end{aligned}$$

$$\implies \hat{\theta}_{\text{MLE}} = \arg_{\theta} l'(\theta) = 0$$